Surprising New Insights on the Early Origins of Autism: Implications for Support from Birth through Adulthood

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SCHOOL OF MEDICINE

Disclosure of Interests



- Industry Consulting: Cognoa
- Stock Equity: None
- Royalties: Western Psychological Services The Social Responsiveness Scale (SRS-2)
- Research Support:
 - NICHD
 - NIMH
 - U.S. Centers for Disease Control

Rare Germline Mutations



Figure 2 | **Diagnostic yield for ID over time.** Graphical overview of the diagnostic yield for moderate to severe intellectual disability (ID) (excluding Down syndrome, which represents 6–8% of all ID) over time. Solid line indicates the mean of published studies, and the shaded background indicates the lower and upper boundaries of reported diagnostic yields. In the 1970s, conventional karyotyping became a routine diagnostic test and provided a conclusive diagnosis in 3–6.5% of ID cases. The

Insufficient Evidence for "Autism-Specific" Genes

Scott M. Myers,^{1,*} Thomas D. Challman,¹ Raphael Bernier,² Thomas Bourgeron,³ Wendy K. Chung,^{4,5} John N. Constantino,^{6,7} Evan E. Eichler,⁸ Sebastien Jacquemont,⁹ David T. Miller,¹⁰ Kevin J. Mitchell,^{11,12} Huda Y. Zoghbi,^{13,14,15,16,17} Christa Lese Martin,¹ and David H. Ledbetter^{1,*}

Despite evidence that deleterious variants in the same genes are implicated across multiple neurodevelopmental and neuropsychiatric disorders, there has been considerable interest in identifying genes that, when mutated, confer risk that is largely specific for autism spectrum disorder (ASD). Here, we review the findings and limitations of recent efforts to identify relatively "autism-specific" genes, efforts which focus on rare variants of large effect size that are thought to account for the observed phenotypes. We present a divergent interpretation of published evidence; discuss practical and theoretical issues related to studying the relationships between rare, large-effect deleterious variants and neurodevelopmental phenotypes; and describe potential future directions of this research. We argue that there is currently insufficient evidence to establish meaningful ASD specificity of any genes based on large-effect rare-variant data.

- --association with intellectual disability
- --timing of effect
- --absence of sex differences



Recurrence Rates and Inherited Transmission in Autism

| MZ concordance: | 90% |
|--------------------------|-----|
| DZ concordance: | 20% |
| Non-twin sib recurrence: | 18% |
| Half-sib recurrence | 6% |
| General population risk | 1% |

Table. Autism Spectrum Disorder Heritability Model Comparisons and Parameter Estimates

| | Model Comparison Measures | | | | Estimated Variance (95% CI) ^a | | | | | | |
|---------------------|---------------------------|---------|-------|----------------------|--|------------------|------------------|------------------|-------------------------------|--|--|
| | No. of Model | | Diff | | Additive Genetic (Narrow-Sense | Nonadditive | Environment | | Total Genetic (Broad-Sense | | |
| Models ^b | Parameters | -2LL | - 2LL | P Value ^c | Heritability) | Genetic | Shared | Nonshared | Heritability) | | |
| ACDE | 14 | 146 836 | NA | NA | 0.69 (0.40-0.86) | 0.10 (0.00-0.38) | 0.04 (0.00-0.14) | 0.16 (0.05-0.30) | 0.80 (0.59-0.95) | | |
| ACE | 13 | 146 836 | 0.4 | .52 | 0.77 (0.58-0.87) | NA | 0.03 (0.00-0.13) | 0.20 (0.13-0.30) | 0.77 (0.58-0.87) | | |
| ADE | 13 | 146 836 | 0.8 | .38 | 0.80 (0.68-0.87) | 0.05 (0.00-0.26) | NA | 0.15 (0.05-0.21) | 0.85 (0.79-0.95) | | |
| CDE | 13 | 146 856 | 20.9 | <.001 | NA | 0.64 (0.48-0.75) | 0.25 (0.21-0.29) | 0.11 (0.03-0.24) | 0.64 (0.48-0.75) | | |
| AE | 12 | 146 836 | 0.9 | .64 | 0.83 (0.79-0.87) | NA | NA | 0.17 (0.13-0.21) | 0.83 (0.79-0.87) | | |
| DE | 12 | 147 100 | 264 | <.001 | NA | 0.99 (0.97-1.00) | NA | 0.01 (0.00-0.03) | 0.99 (0.97-1.00) | | |
| CE | 12 | 146 897 | 61 | <.001 | NA | NA | 0.39 (0.37-0.41) | 0.61 (0.59-0.63) | NA | | |
| E | 11 | 147 996 | 1160 | <.001 | NA | NA | NA | 1.00 (1.00-1.00) | NA | | |

Abbreviations: 2LL, $2 \times \text{logarithm}$ of the likelihood; Diff – 2LL, $2 \times \text{difference}$ in log-likelihood between the model and the full model; NA, not applicable.

^a The 95% CIs are 2-sided CIs. Variances are based on the tetrachoric correlations. The unadjusted tetrachoric correlation (SD) was estimated to 0.87 (0.08) and 0.40 (0.10) for monozygotic and dizygotic twins; 0.41 (0.01) for full siblings; 0.22 (0.03) and 0.17 (0.04) for maternal and paternal half siblings.

^b All models adjusted for sex and birth cohort. The genetic terms for each model

are shown in each row, which include additive genetic effect (A; inherited additive effects of different alleles), shared environmental factors (C; nongenetic influences contributing to similarity within sibling pairs), nonadditive (dominant) genetic factors (D; interaction effects between alleles at the same locus), and nonshared environmental factors (E; making siblings dissimilar).

^c *P* value for testing the hypothesis: the parameters not in the model but in the full model are all equal to 0.

Quantitative autistic traits

Tracing phenotypic manifestations of inherited variant liability in families







Male First Degree Relatives: Shifted Unimodal Distribution



www.thelancet.com/neurology 2015









THE JOURNAL OF CHILD PSYCHOLOGY AND PSYCHIATRY

Journal of Child Psychology and Psychiatry *:* (2012), pp **-**

Vanessa Hus,¹ Somer Bishop,² Katherine Gotham,¹ Marisela Huerta,³ and Catherine Lord³

¹Department of Psychology, University of Michigan; ²Cincinnati Children's Hospital Medical Center; ³Department of Psychiatry, Weill-Cornell Medical School

Probands Siblings Males (n = 2,056)Females (n = 312)Females (n = 1,023)Males (n = 890)9.49 (3.71623) Age (years) 8.74 (3.32) 8.90 (3.60) 9.46 (3.65) SRS-Raw 97.56 (26.82) 99.32 (27.24) 20.53 (15.44) 17.22 (13.02) SRS T-score 43.70 (7.39) 80.56 (12.83) 89.63 (15.05) 44.26 (7.19) VSOC 71.54 (12.57) 70.15 (12.71) 101.77 (11.93) 103.19 (11.29) CBCL-E 56.39 (10.7) 57.79 (10.2) 46.84 (9.81) 46.19 (9.42) 60.35 (9.47) 59.96 (9.98) 48.29 (10.19) 47.32 (9.96) CBCL-I VEC 10.23 (3.06) 9.71 (3.03) 16.02 (2.37) 16.39 (2.34) ADI-Current 17.04 (7.30) 17.59 (7.80) ADOS-CSS 7.43 (1.68) 7.43 (1.73) 85.81 (25.70) 78.12 (25.18) NVIQ

Parent SRS reports in the Simons Simplex Collection

3 S.D. difference between children with ASD (probands) and their unaffected sibs Trait correlation with Attention Problems on the order of r=0.50



Annu. Rev. Clin. Psychol. 2021. 17:X–X https://doi.org/10.1146/annurev-clinpsy-081219-110503 Copyright © 2021 by Annual Reviews.

Constantino • Charman • Jones www.annualreviews.org • A Developmental Substructure for Autism

1. A developmental sub structure for autism

"Grand Unified Theory" ^(C) Inheritance exerts effects via <u>independent</u> components of additive genetic liability, MOST NOT SPECIFIC TO AUTISM, each with its own link to early developmental endophenotypes which jointly give rise to the autistic syndrome



Circulation: Overall Regulation A C Guyton, et al. Annual Review of Physiology 1972 34:1, 13-44

What exactly is heritable in the causation of autism?



Figure 3. Salience maps for typically-developing 2-year-olds (top) and for 2-year-olds with au tism (bottom). Images at right show color data scaled from black to transparent and overlaid on the still image from the video scene.

Courtesy, Ami Klin and Warren Jones, Emory University

Social visual engagement in infants is under stringent genetic control



Constantino et al., Nature (2017)







- 1. Most children with ASD are in the lower range of eye and mouth looking
- 2. Some children in the lower range of eye and mouth looking are typically-developing

Early eye gaze abnormalities as "necessary but not sufficient" to cause ASD

Psychological Medicine

cambridge.org/psm

Original Article

Attention-deficit/hyperactivity disorder and risk for psychiatric and neurodevelopmental disorders in siblings

Elina Jokiranta-Olkoniemi¹, Keely Cheslack-Postava², Petteri Joelsson¹, Auli Suominen¹, Alan S. Brown^{2,3} and Andre Sourander^{1,2}

Methods. Every child born in Finland in 1991–2005 and diagnosed with ADHD in 1995–2011 were identified from national registers. Each case was matched with four controls on sex, place, and date of birth. The full siblings of the cases and controls were born in 1981–2007 and diagnosed in 1981–2013. In total, 7369 cases with 12 565 siblings and 23 181 controls with 42 753 siblings were included in the analyses conducted using generalized estimating equations.

Table 2. Associations between ADHD and psychiatric and neurodevelopmental disorders among the siblings of cases and matched controls

| | ADHD ^a | | Adjusted ^b (model I) | | Adjusted ^c (model II) | | Adjusted ^d (model III) | |
|------------------------------------|----------------------|---------------------------|---------------------------------|--------------|----------------------------------|--------------|-----------------------------------|--------------|
| | Case n = 7369 (%) | Control n = 23 181 (%) | RR | (95% CI) | RR | (95% CI) | RR | (95% CI) |
| ADHD | 1134 (15.4) | 598 (2.6) | 7.1 | (6.4-7.8)*** | 6.5 | (5.9–7.2)*** | 5.7 | (5.1-6.3)*** |
| Conduct and oppositional disorders | 705 (9.6) | 461 (2.0) | 5.6 | (5.0-6.3)*** | 5.0 | (4.5–5.7)*** | 4.0 | (3.5–4.5)*** |
| ASD | 322 (4.4) | 264 (1.1) | 4.6 | (3.9-5.4)*** | 4.4 | (3.7–5.2)*** | 3.9 | (3.3-4.6)*** |

Mous et al. Journal of Neurodevelopmental Disorders (2017) 9:32 DOI 10.1186/s11689-017-9212-y

Candidate #4 ADHD

RESEARCH

Attention and motor deficits index non-specific background liabilities that predict autism <u>recurrence</u> in siblings

Sabine E. Mous^{1,2}, Allan Jiang², Arpana Agrawal² and John N. Constantino^{2*}

| | | | - | | | - | | | | |
|--|---------|-------|---------|-------|---------|--------|---------|--------|---------|--------|
| | Model 1 | | Model 2 | | Model 3 | | Model 4 | | Model 5 | |
| | β | р | β | р | β | р | β | р | β | р |
| Proband SRS-2 score (teacher-report) | 0.30 | 0.086 | 0.26 | 0.098 | 0.19 | 0.109 | 0.19 | 0.111 | | |
| Sibling TRF ADHP score (teacher-report) | | | 0.45 | 0.005 | 0.24 | 0.066 | 0.25 | 0.074 | 0.24 | 0.063 |
| Sibling DCDQ score (parent-report) | | | | | -0.60 | <0.001 | -0.60 | <0.001 | -0.62 | <0.001 |
| TRF ADHP x DCDQ interaction | | | | | | | 0.03 | 0.793 | | |
| Adjusted R ² | 0.0 | 059 | 0.2 | 247 | 0.5 | 554 | 0.5 | 540 | 0.5 | 530 |

Table 4. Linear regression analyses predicting parent-reported autistic trait severity in siblings

S. Mous *Erasmus University* The Netherlands





ARTICLE

Open Access

Behavioral predictors of autism recurrence are genetically independent and influence social reciprocity: evidence that polygenic ASD risk is mediated by separable elements of developmental liability

Alexa Pohl¹, Warren R. Jones², Natasha Marrus³, Yi Zhang³, Ami Klin² and John N. Constantino⁴

Table 3 Results of linear regression analysis examining the joint contribution of three behavioral predictors of autismrecurrence (measured at 36-48 months) to variation in autism-related variation in early childhood reciprocal socialbehavior

| Outcome modeled | Adj R ² | BPAR | В | t | Sig | Δ Adj R ² |
|------------------|--------------------|-------------------------------------|--------|--------|---------|----------------------|
| SRS at 36 months | 0.35 | Biparental QAT | 0.255 | 5.568 | < 0.001 | 0.06 |
| | | Variation in attentional impairment | 0.355 | 3.996 | <0.001 | 0.12 |
| | | Variation in motor coordination | -0.242 | -3.830 | <0.001 | 0.05 |
| | | Site | .079 | 1.252 | 0.212 | 0.00 |

Adjusted R square is reported for the full regression model, along with changes in adjusted R square that occur when a given individual behavioral trait is excluded from the model, and the result compared with that for the full model. A companion table for SRS outcome at 48 months (for which there were fewer twin pairs with complete data) yielded highly comparable results and is provided in Supplementary Table 2 SRS social responsiveness scale

NEUROSCIENCE —

RESEARCH ORGANIZATION

L. L. Orefice/Neuroscience 445 (2020) 120-129

Peripheral Somatosensory Neuron Dysfunction: Emerging Roles in Autism Spectrum Disorders

Lauren L. Orefice

Department of Molecular Biology, Massachusetts General Hospital and Department of Genetics, Harvard Medical School, 185 Cambridge Street, Boston, MA 02114, USA

Tactile sensitivity \rightarrow acquired deviation in brain / behavioral development



An Analogy...









Autism Spectrum Disorder



ARTICLE



∂ OPEN ACCESS

Check for updates

Deconstructing autism: from unitary syndrome to contributory developmental endophenotypes

John N. Constantino 🝺

Departments of Psychiatry and Pediatrics, Washington University School of Medicine, St Louis, MO, USA

Deconstructing familial autism

(by early developmental endophenotypes)



COMMENTARY

Open Access



New guidance to seekers of autism biomarkers: an update from studies of identical twins

John N. Constantino^{*} 🖸

Abstract

Background: The autism spectrum disorders (ASD) are common neuropsychiatric conditions of childhood for which the vast proportion of population risk is attributable to inheritance, and for which there exist few if any replicated biomarkers.

Main body: This commentary summarizes a set of recent studies involving identical (monozygotic, MZ) twins which, taken together, have significant implications for the search for biomarkers of inherited susceptibility to autism. A first is that variation-in-severity of the condition (above the threshold for clinical diagnosis) appears more strongly influenced by stochastic/non-shared environmental influences than by heredity. Second is that there exist disparate early behavioral predictors of the familial recurrence of autism, which are themselves strongly genetically influenced but largely independent from one another. The nature of these postnatal predictors is that they are trait-like, continuously distributed in the general population, and largely independent from variation in general cognition, thereby reflecting a developmental substructure for familial autism. A corollary of these findings is that autism may arise as a developmental *consequence* of an allostatic load of earlier-occurring liabilities, indexed by early behavioral endophenotypes, in varying permutations and combinations. The clinical threshold can be viewed as a "tipping point" at which stochastic influences and/or other non-shared environmental influences assert much stronger influence on variation-inseverity (a) than do the genetic factors which contributed to the condition in the first place, and (b) than is observed in typical development.

Conclusion: Biomarkers identified on the basis of association with clinical symptom severity in ASD may reflect *effects* rather than *causes* of autism. The search for biomarkers of pathogenesis may benefit from a greater focus on traits that predict autism recurrence, among both clinical and general populations. In case–control studies, salient developmental liabilities should be systematically measured in both cases and controls, to avoid the erosion in statistical power (i.e., to detect differences) that can occur if control subjects carry sub-clinical aggregations of the same unmeasured traits that exert causal influences on the development of autism.

2. Stochastic Influences on Severity

Differential Heritability Below versus Above the Clinical Threshold for Severity

ORIGINAL RESEARCH

20

0

25

50



On the Nature of Monozygotic Twin Concordance and Discordance for Autistic Trait Severity: A Quantitative Analysis



100

125

is approached and exceeded

*****NO age effects observed

SRS of Higher-Scoring MZ Twin

75

Castelbaum et al., Behav Genet 2019

ADOS SCORES OF IDENTICAL CO-TWINS IN AGRE



N. B. Acutal MZ twin discordance (4%) is much rarer than presumed by families / community



NIH Public Access

Author Manuscript

Res Autism Spectr Disord. Author manuscript; available in PMC 2009 August 28.

Published in final edited form as:

Res Autism Spectr Disord. 2008 April 1; 2(2): 320-331. doi:10.1016/j.rasd.2007.08.002.

Genetic and Environmental Influences on Symptom Domains in Twins and Siblings with Autism

Carla A. Mazefsky, Ph.D, Robin P. Goin-Kochel, P.hD, Brien P. Riley, P.hD, Hermine H. Maes, Ph.D, and The Autism Genetic Resource Exchange Consortium^{*}

N=1,294 twins and siblings with ASD, AGRE ADI-R Data

Unstandardized Estimates of Additive Genetic (A), Dominant Genetic (D), and Unique Environmental (E) Influences for ADI Nonverbal Communication and Social Dysfunction Total Scores, With 95% Confidence Intervals in Parentheses

| Nonverbal Communication | | | Social Dys | function | Common influences on Both Phenotypes | | | |
|--|--------------------------------|----------------------------------|-----------------|-----------------|--------------------------------------|-----------------|--|--|
| Estimates from the full ADE model, with males and females free | | | | | | | | |
| | Males | Females | Males | Females | Males | Females | | |
| А | .00 (.00 – .29) | .08 (.00 – .53) | .20 (.09 – .35) | .08 (.00 – .53) | .00 (.00 – .15) | .06 (.00 – .39) | | |
| D | .44 (.25 – .58) | .58 (.07 – 1.02) | .16 (.00 – .34) | .38 (.03 – .96) | .26 (.04 – .43) | .46 (.05 – .75) | | |
| Е | .48 (.35 – .71) | .48 (.32 – .71) | .57 (.39 – .77) | .76 (.35 – .98) | .36 (.20 – .51) | .46 (.2580) | | |
| Estimates from the | he best fitting model, with ma | les and females equal for all pa | rameters | | | | | |
| А | .01 (.00 – .19) | | .16 (.00 – .35) | | 02 (0615) | | | |
| D | .46 (.13 – .64) | | .24 (.01 – .56) | | .33 (.04 – .50) | | | |
| Е | .51 (.35 – .71) | | .59 (.4080) | | .39 (.24 – .57) | | | |

3. The Nature of Sex Effects...

...may not be what we have thought they are; two implications of the Female Protective Effect (FPE):

a.) The Carter Effect

b.) Risk of ASD among the Children of Sisters of ASD Probands...





Inherited Risk for Autism Through Maternal and Paternal Lineage

Dan Bai, Natasha Marrus, Benjamin Hon Kei Yip, Abraham Reichenberg, John N. Constantino, and Sven Sandin



Increasing genetic liability for ASDs

Figure 1. The expected autism spectrum disorder (ASD) liability under the female protective effect. Example: an individual has ASD and, hypothetically, he has 10 sisters (left panel, circles) and 10 brothers (right panel, squares). Under a female protective effect, 1) only the sisters with very high genetic liability will be diagnosed with ASD; and 2) for the brothers, there will be more diagnosed cases of ASD: in this example, a 3:1 male-to-female ratio. In the next generation (not shown in the pedigree), because it is more likely that undiagnosed sisters carry a moderate-to-high genetic liability, the relative risk of ASD among offspring to the undiagnosed siblings is expected to be higher among those to the sisters than among those to the brothers.

Table 2. Relative Risks for ASD Among Participants With ASD-Affected Uncle(s)/Aunt(s) Compared With Participants With Uncle(s)/Aunt(s) Free From ASD Diagnosis

| | | Persor Foll | n-Years of ow-up | Rates c 100,000 F | of ASD per Person-Years | Re | elative Risk (95% C | Risk (95% Cl) ^a | | |
|---|----------------------------------|----------------|---------------------|----------------------|----------------------------|------------------|-------------------------|----------------------------|--|--|
| E | Exposure | Exposed | Unexposed | Exposed | Unexposed | Crude | Adjusted 1 ^b | Adjusted 2 ^c | | |
| I | Maternal Lineage | | | | | | | | | |
| | Uncle(s) affected by ASD | 19,662 | 6,675,534 | 401.79 | 151.07 | 3.22 (2.54-3.91) | 2.75 (2.15–3.36) | 1.92 (1.49–2.37) | | |
| | Aunt(s) affected by ASD | 9984 | 6,337,615 | 340.54 | 152.71 | 2.73 (1.88–3.73) | 2.36 (1.62-3.21) | 1.73 (1.19–2.35) | | |
| | Uncle(s)/aunt(s) affected by ASD | 29,646 | 13,013,149 | 381.16 | 151.87 | 3.05 (2.52-3.64) | 2.62 (2.17-3.14) | 1.88 (1.54–2.26) | | |
| F | Paternal Lineage | | | | | | | | | |
| | Uncle(s) affected by ASD | 13,274 | 6,723,265 | 301.35 | 153.23 | 2.39 (1.72-3.13) | 2.09 (1.50-2.74) | 1.63 (1.16–2.16) | | |
| | Aunt(s) affected by ASD | 7342 | 6,292,408 | 190.68 | 151.99 | 1.52 (0.79–2.30) | 1.33 (0.69–2.01) | 1.07 (0.56–1.63) | | |
| | Uncle(s)/aunt(s) affected by ASD | 20,616 | 13,015,674 | 261.94 | 152.63 | 2.08 (1.53-2.67) | 1.82 (1.32-2.34) | 1.44 (1.05–1.86) | | |

"Affected by ASD" refers to an individual with a clinical diagnosis of ASD.

ASD, autism spectrum disorder; CI, confidence interval.

^aBootstrapped 95% CI: 2.5%–97.5% percentiles of estimates from 1000 bootstrapped samples.

^bAdjusted for the birth year of the participant, the mother, and the uncle/aunt.

^cAdjusted for covariates in Adjusted 1 and any mental illness (yes/no) of the mother and the uncle/aunt.

First epidemiologic estimate of risk to second generation offspring:

For sisters of an ASD-affected individual, risk of autism = $2-3 \times 2-3 \times 2-5 \times 2-5$

4. Shift Happens

Background genetic variation within the normal range shifts the "starting point" for the deleterious effects of some (not all) mutations Research

Original Investigation

The Role of Parental Cognitive, Behavioral, and Motor Profiles in Clinical Variability in Individuals With Chromosome 16p11.2 Deletions

VARIABLE PENETRANCE

Andres Moreno-De-Luca, MD; David W. Evans, PhD; K. B. Boomer, PhD; Ellen Hanson, PhD; Raphael Bernier, PhD; Robin P, Goin-Kochel, PhD; Scott M. Myers, MD; Thomas D. Challman, MD; Daniel Moreno-De-Luca, M Mylissa M. Slane, MS; Abby E. Hare, PhD; Wendy K. Chung, MD; John E. Spiro, PhD; W. Andrew Faucet Orista L. Martin, PhD; David H. Ledbetter, PhD

Shift happens: family background influences clinical variability in genetic neurodevelopmental disorders

Brenda Finucane, MS, LGC¹, Thomas D. Challman, MD, FAAP¹, Christa Lese Martin, PhD, FACMG¹ and David H. Ledbetter, PhD, FACMG¹





Conclusions

- The strong family genetic liability to *autism* operates via multiple independent liabilities, permutations and combinations of which give rise to autism; in this sense autism is "fractionable" before it develops BUT NOT AFTER
- Autism severity is strongly influenced by stochastic influences / nonshared environmental effects—this is true above, but not below, the clinical threshold
- If the components of the developmental sub structure for autism are not measured among CONTROL subjects (or as indexed by parental background traits in "shift" studies), expect erosion in *statistical power* to identify a genotype-phenotype association (or any biomarker association related to cause of ASD
- Framing the ASD sex bias as a function of *male sensitivity* rather than female protection has major implications for the design of future studies (including cell, organoid, rodent) of the effects of sex on ASD
- New approach to *early behavioral intervention*?: "nudging" endophenotypic liabilities into the safe/normative range <u>before</u> autism develops
 - Social visual disengagement
 - Inattention
 - Motor coordination deficits (cerebellar learning)
 - ? Tactile sensitivity

REVIEW

Multiplexed assays of variant ef genotype-phenotype atlas

Jochen Weile^{1,2,3,4} · Frederick P. Roth^{1,2,3,4}



Fig. 1 Percentage of variants of uncertain significance (VUS) among missense allele Clinvar records over time from 1990 until 2017



NATIONAL BRAIN-GENE REGISTRY Nitional Center rAdvancing Cranstational Sciences

Co-registration of

- A) clinically-identified variants of unknown significance (VUS)
- B) Rapid Standardized (Virtual) Neurobehavioral Phenotyping Data
- C) Electronic Health Record Data



FIGURE 2. Most recent intelligence quotient score as of age 8 years among children with autism spectrum disorder for whom test data were available, by sex and race/ethnicity - Autism and Developmental Disabilities Monitoring Network, nine sites,* United States, 2014

2018

April 27, 2018



Abbreviations: ASD = autism spectrum disorder; F = female; IQ = intelligence quotient; M = male.

* Includes nine sites (Arizona, Arkansas, Colorado, Georgia, Maryland, Minnesota, New Jersey, North Carolina, and Tennessee) that had intellectual ability data available for \geq 70 of children who met the ASD case definition (n = 3,714).

Maenner et al. MMWR 2020

Prevalence estimates were approximately identical for non-Hispanic white and non-Hispanic black children (18.5 and 18.3, per thousand, respectively)

Among children with ASD for whom data on intellectual or cognitive functioning were available, 33% were classified as having intellectual disability (intelligence quotient [IQ] \leq 70); this percentage was higher among black and Hispanic than white children (47%, 36%, and 27%, respectively).

Black children with ASD were less likely to have a first evaluation by age 36 months than were white children with ASD (40% versus 45%).

Black children with IQ ≤70 had a later median age at ASD diagnosis than white children with $IQ \leq 70$ (48 months versus 42 months).

PEDIATRICS

Structural Racism and Autism

Sarabeth Broder-Fingert, MD, MPH,^a Camilla Mateo, MD, MPH,^b Katherine E. Zuckerman, MD, MPH^c

- We believe these data shed light on structural racism as a driver of inequity for children with ASD.
- 31.3% of parents cited availability of professionals as a barrier to diagnosis.
 - Workforce capacity has long been a major challenge in the field of ASD.
 - This issue may be compounded for AA families who may be more likely than white families to live in areas with few ASD diagnostic specialists and also be more likely to rely on Medicaid.
 - Medicaid may aggravate inequities because low rates of Medicaid reimbursement for diagnostic services can create challenges to supporting and maintaining a workforce.
- Thus, AA families may be suffering from a lack of available workforce related to <u>where</u> <u>they live</u> (as a result of structural racism in housing policy), whereas clinics serving these communities may be <u>unable to expand capacity</u> (because of structural racism in Medicaid reimbursement rates).
- These issues are likely exacerbated by additional issues on the provider level including a lack of cultural humility and workforce diversity. Notably, a recent workforce survey showed that only 2% of developmental-behavioral or neurodevelopmental pediatricians are AA.

THANK YOU

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Genetic relationships between Mendelian and Complex Diseases



Blair et al. Cell 2013

ORIGINAL RESEARCH



On the Nature of Monozygotic Twin Concordance and Discordan for Autistic Trait Severity: A Quantitative Analysis

Lauren Castelbaum¹ · Chad M. Sylvester¹ · Yi Zhang¹ · Qiongru Yu¹ · John N. Constantino^{1,2}

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General Population MZ Twins, n=268 pairs



Clinically-ascertained MZ twins, n = 79 *pairs*





doi:10.1038/nature22999

Infant viewing of social scenes is under genetic control and is atypical in autism

John N. Constantino^{1,2,3}, Stefanie Kennon-McGill¹, Claire Weichselbaum¹, Natasha Marrus^{1,3}, Alyzeh Haider¹, Anne L. Glowinski¹, Scott Gillespie⁴, Cheryl Klaiman^{5,6}, Ami Klin^{5,6,7} & Warren Jones^{5,6,7}

Table 1 Falconer's heritability and twin-twin correlations for two behavioral predictors of ASD recurrence (Attention Problems and variation in Motor Coordination) and autism-related variation in reciprocal social behavior in the fourth year of life

| | Attention problem (CBCL) | Motor coordination (Little DCDQ) | Reciprocal social behavior (SRS-2) |
|---------------------------|--------------------------|----------------------------------|------------------------------------|
| MZM twin-twin correlation | .527 | .937 | .784 |
| DZM twin-twin correlation | 013 | .734 | .327 |
| MZF twin-twin correlation | .583 | .898 | .913 |
| DZF twin-twin correlation | 081 | .625 | .352 |
| Heritability (males) | .53* | ,41 | .91 |
| Heritability (females) | .58* | .55 | .91* |

The analysis excluded opposite sex twin pairs and pairs with uncertain zygosity; they are derived from 66 MZ pairs (34 male-male and 32 female-female) and 58 DZ pairs (30 male-male, 28 female-female) with complete data for these three variables

*For heritability estimate >1, MZ concordance rate is used by convention

MZM monozygotic male

DZM dizygotic male

MZF monozygotic female

DZF dizygotic female

CBCL child behavior checklist

Little DCDQ developmental coordination disorder questionnaire

SRS-2 social responsiveness scale, 36 months

| Table 2 | Matrix depicting within-individual correlations (Pearson's r) between autism recurrence predictors in 174 | |
|-----------|---|--|
| general p | population twins (one twin selected at random from each pair) | |

| | Attentional impairment | QAT-p | Motor coord | SRS 36 months | SRS 48 months |
|------------------------------------|------------------------|-------|-------------|---------------|---------------|
| Attentional impairment | 1 | | | | |
| QAT-p | .20 | 1 | | | |
| Motor coordination | 20 | 16 | 1 | | |
| SRS-2 (36 months) | .46 | .38 | 36 | 1 | |
| SRS-2 (48 months., <i>n</i> = 160) | .48 | .28 | 35 | .71 | 1 |

All correlations above 0.29 were statistically significant at p < .01; none of the bivariate associations between BPARs reached this threshold. Bivariate cross-twin cross-trait correlations encompassing attentional impairment, QAT-p, and motor coordination were uniformly non-statistically significant QAT-p Quantitative Autistic Traits of Parents, as measured by the Social Responsiveness Scale, Adult Version

Points of bifurcation in Waddington's epigenetic landscape

